
Probability and Statistics

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Topics

- 1 Large Sample Test for a General Population Mean μ
- 2 Test for a Normal Mean μ of When σ is Unknown

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Objectives

Objectives:

- Carry out a hypothesis test for a general population mean μ when the sample size is large.
- Carry out a hypothesis test for a normal mean μ when the standard deviation σ is unknown.

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Large Sample Test for a General Population Mean μ

(8.2)

- When the **sample size n is large**, two things are true:
 1. Regardless of the shape of the population, by the Central Limit Theorem,

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1) \quad (\text{approximately})$$

2. The **sample standard deviation S** will remain fairly **constant** from one sample to the next, and **approximately equal to σ** .

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- As a consequence, **when n is large**,

$$\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim N(0, 1) \quad (\text{approximately})$$

even if the sample is from **non-normal** population.

(Note that σ was replaced by S above.)

- It follows that **if H_0 is true** (so $\mu = \mu_0$),

$$\frac{\bar{X} - \mu_0}{S/\sqrt{n}} \sim N(0, 1) \quad (\text{approximately})$$

- Thus we can use this as our **test statistic** in a **one-sample z test for μ** .

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One-Sample Z Test Statistic:

$$Z = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

When H_0 is true, $Z \sim N(0, 1)$.

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Sampling Distribution of the Test Statistic Under H_0 :

If Z is the one-sample Z test statistic (from the previous slide), then when

$$H_0 : \mu = \mu_0$$

is true,

$$Z \sim N(0, 1).$$

- The **p -value** is the appropriate tail area under the $N(0, 1)$ curve. (See Slides 17.)
- In practice, n is **large enough** if $n > 40$.

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One-Sample Z Test for μ when n is Large

Assumptions: The data x_1, x_2, \dots, x_n are a random sample from any distribution whose mean is μ and n is large.

Null hypothesis: $H_0 : \mu = \mu_0$.

Test statistic value: $z = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$.

Decision rule: Reject H_0 if $p\text{-value} < \alpha$.

Alternative hypothesis

$H_a : \mu > \mu_0$

$H_a : \mu < \mu_0$

$H_a : \mu \neq \mu_0$

P-value = area under $N(0, 1)$ distribution:

to the right of z

to the left of z

to the left of $-|z|$ and right of $|z|$

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Test for a Normal Mean μ of When σ is Unknown (8.3)

- Recall (Slides 15) that if X_1, X_2, \dots, X_n are a random sample from a **normal** population, the random variable

$$\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1),$$

a **t distribution** with $n - 1$ **degrees of freedom**.

- It follows that if H_0 is true (so $\mu = \mu_0$),

$$\frac{\bar{X} - \mu_0}{S/\sqrt{n}} \sim t(n-1).$$

- Thus we can use this as our **test statistic** in a **one-sample t test for μ** .

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- The **test statistic** for the **one-sample t test for μ** is

One-Sample t Test Statistic:

$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

When H_0 is true, $T \sim t(n-1)$.

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Sampling Distribution of the Test Statistic Under H_0 :

If T is the one-sample t test statistic, then when

$$H_0 : \mu = \mu_0$$

is true,

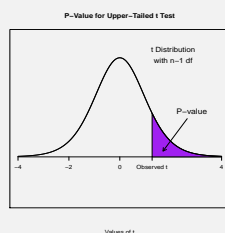
$$T \sim t(n-1).$$

- The **p-value** is the appropriate tail area under the $t(n-1)$ curve. (See the next few slides.)

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P-Value: For the **one-sample t test**, the **p-value** is the tail area under the $t(n-1)$ curve:

- To the **right** of the observed t if the alternative hypothesis is $H_a : \mu > \mu_0$:



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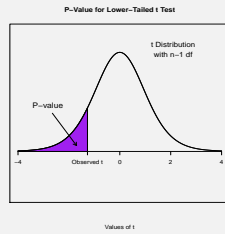
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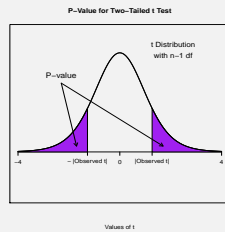
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1. To the **left** of the observed t if the alternative hypothesis is $H_a : \mu < \mu_0$:



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1. To the **left** of $-|t|$ **and right** of $|t|$ if the alternative hypothesis is $H_a : \mu \neq \mu_0$:



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One-Sample t Test for μ

Assumptions: The data x_1, x_2, \dots, x_n are a random sample from a *normal* distribution.

Null hypothesis: $H_0 : \mu = \mu_0$.

Test statistic value: $t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$.

Decision rule: Reject H_0 if $p\text{-value} < \alpha$.

Alternative hypothesis	P-value = area under the t distribution with $n - 1$ d.f.:
$H_a : \mu > \mu_0$	to the right of t
$H_a : \mu < \mu_0$	to the left of t
$H_a : \mu \neq \mu_0$	to the left of $- t $ and right of $ t $

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Example

A quality control engineer monitors a machine that puts cereal into boxes.

According to the label, each box is supposed to contain **16 oz** of cereal.

The machine will need to be adjusted if the boxes are systematically being **under-filled** or **over-filled**.

From past experience, the engineer knows that the weight (ounces) of the cereal in a box follows a **normal** distribution.

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To decide if the boxes are being **under-filled or overfilled**, the engineer will test the **hypotheses**

$$H_0 : \mu = 16$$

$$H_a : \mu \neq 16$$

where μ is the true (unknown) population mean weight.

A random sample of **ten** boxes gives

$$\bar{x} = 16.6 \quad \text{and} \quad s = 0.9.$$

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Thus the observed **test statistic** is

$$\begin{aligned} t &= \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \\ &= \frac{16.6 - 16}{0.9/\sqrt{10}} \\ &= 2.11. \end{aligned}$$

Thus the **sample mean** weight, $\bar{x} = 16.6$, is about **2.11 standard errors above 16** ounces.

The **p-value** is the **probability** that we'd get a t value this far away from zero (in either direction) by chance **if the population mean μ was 16**.

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From the **two tail** areas of the **sampling distribution** that the test statistic would follow under H_0 (the $t(9)$ distribution), to the **right of 2.11 and left of -2.11**,

$$\text{p-value} = 2(0.033) = 0.066.$$

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Thus we'd get a result like the one we got **6.6%** of the time **even if the population mean μ was 16** ounces.

Using a **level of significance** $\alpha = 0.05$, the **decision rule** is

$$\begin{aligned} &\text{Reject } H_0 \text{ if p-value} < 0.05. \\ &\text{Fail to reject } H_0 \text{ if p-value} \geq 0.05. \end{aligned}$$

Because $0.066 \geq 0.05$, we **fail to reject H_0** .

There's **no statistically significant evidence** that the population mean cereal box weight μ is different from 16 ounces.

The result that the engineer got (by taking a random sample) can be explained by chance variation (sampling error).

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