

Statistical Methods

Nels Grevstad

Metropolitan State University of Denver

ngrevsta@msudenver.edu

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Topics

- 1 Two-Factor ANOVA with More Than One Observation Per Cell

Objectives

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- State the treatment effects version of the two-factor ANOVA model when $K > 1$.
- Carry out two-factor ANOVA F tests for the interaction effect and main effects of Factors A and B when $K > 1$.
- Interpret the interaction effect and main effects in the two-factor ANOVA model.

Two-Factor ANOVA with $K \geq 2$

- Two-factor studies usually involve more than one observation per group.

Example

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Twenty-four programmers, classified according to **type** and **years of experience**, were asked to predict the number of programmer-days required to complete a large project about to be initiated.

The table below shows their prediction errors (actual minus predicted programmer-days).

		Factor B: Years of Experience			
		Less than 5 ($j = 1$)	5 to 10 ($j = 2$)	More than 10 ($j = 3$)	
Factor A: Type of Experience	Small Systems Only ($i = 1$)	240	110	56	$\bar{X}_{1..} = 129.7$
		206	118	60	
		217	103	68	
		225	95	58	
	Small and Large Systems ($i = 2$)	71	47	37	$\bar{X}_{2..} = 48.6$
		53	52	33	
68		31	40		
		57	49	45	
		$\bar{X}_{.1.} = 142.1$	$\bar{X}_{.2.} = 75.6$	$\bar{X}_{.3.} = 49.6$	$\bar{X}_{...} = 89.1$

Here's a different version of the table showing the so-called **group means**.

		Factor B: Years of Experience			
		Less than 5 ($j = 1$)	5 to 10 ($j = 2$)	More than 10 ($j = 3$)	
Factor A: Type of Experience	Small Systems Only ($i = 1$)	$\bar{X}_{11.} = 222.0$	$\bar{X}_{12.} = 106.5$	$\bar{X}_{13.} = 60.5$	$\bar{X}_{1..} = 129.7$
	Small and Large Systems ($i = 2$)	$\bar{X}_{21.} = 62.3$	$\bar{X}_{22.} = 44.8$	$\bar{X}_{23.} = 38.8$	$\bar{X}_{2..} = 48.6$
		$\bar{X}_{.1.} = 142.1$	$\bar{X}_{.2.} = 75.6$	$\bar{X}_{.3.} = 49.6$	$\bar{X}_{...} = 89.1$

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- Does **type of experience** effect prediction accuracy (i.e. is there a ***factor A main effect***)?
- Does **years of experience** effect prediction accuracy (i.e. is there a ***factor B main effect***)?
- Is the effect of **years of experience** ***different*** depending on a programmer's **type of experience** (i.e. is there an ***interaction effect***)?

- **Notation:**

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I = The number of levels of Factor A .

J = The number of levels of Factor B .

K = The number of observations (common sample size) in each of the IJ treatment groups.

X_{ijk} = The k th observation at the i th level of Factor A and j th level of Factor B (i.e. in the i, j th group).

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- The data can be laid out in a table as below:

Factor B

		Level $j = 1$	Level $j = 2$...	Level $j = J$	
Factor A	Level $i = 1$	X_{111}	X_{121}	...	X_{1J1}	$\bar{X}_{1..}$
		X_{112}	X_{122}	...	X_{1J2}	
		\vdots	\vdots		\vdots	
		X_{11K}	X_{12K}		X_{1JK}	
	Level $i = 2$	X_{211}	X_{221}	...	X_{2J1}	$\bar{X}_{2..}$
		X_{212}	X_{222}	...	X_{2J2}	
		\vdots	\vdots		\vdots	
		X_{21K}	X_{22K}		X_{2JK}	
	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots
	Level $i = I$	X_{I11}	X_{I21}	...	X_{IJ1}	$\bar{X}_{I..}$
X_{I12}		X_{I22}	...	X_{IJ2}		
\vdots		\vdots		\vdots		
X_{I1K}		X_{I2K}		X_{IJK}		
		$\bar{X}_{.1.}$	$\bar{X}_{.2.}$...	$\bar{X}_{.J.}$	$\bar{X}_{..}$

- (cont'd)

$\bar{X}_{i..}$ = The **Factor A level mean** of all observations at level i of Factor A.

$\bar{X}..j$ = The **Factor B level mean** of all observations at level j of Factor B.

$\bar{X}_{ij.}$ = The **group mean** of the observations in the i, j th group.

$\bar{X}...$ = The **grand mean** of *all* IJK observations.

- When the sample sizes per group are all the same, the grand mean $\bar{X}_{...}$ can be obtained as:
 - The average of the IJ group means.
 - The average of the I Factor A level means.
 - The average of the J Factor B level means.

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- The sample sizes per group **don't** all have to be the same. But we'll only look at the equal-sample size case.
- The data can be **samples** from IJ populations (representing combinations of the levels of the factors) **or** responses to treatments in a **randomized experiment**.

The Two-Factor ANOVA Model

- When $K > 1$, we **no longer** have to assume the effects of the factors are **additive** (which did when $K = 1$ to reduce the number of model parameters that needed to be estimated).

Two-factor ANOVA Model:

$$X_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}, \quad (1)$$

where

μ is a constant called the *true grand mean*.

α_i is the *effect* of the i th level of **Factor A**.

β_j is the *effect* of the j th level of **Factor B**.

γ_{ij} is the *interaction effect* for the i th level of **Factor A** and j th level of **Factor B**.

ϵ_{ijk} are iid $N(0, \sigma)$ *random errors*.

(More formal definitions on the next slide.)

- More formally, let

μ_{ij} = The **population mean** for the i th level of **Factor A** and j th level of **Factor B**.

Then:

$$\mu = \frac{\sum_i \sum_j \mu_{ij}}{IJ},$$

$$\alpha_i = \mu_{i.} - \mu \quad \text{and} \quad \beta_j = \mu_{.j} - \mu,$$

where the **true factor A and B levels means**, $\mu_{i.}$ and $\mu_{.j}$, are

$$\mu_{i.} = \frac{\sum_j \mu_{ij}}{J} \quad \text{and} \quad \mu_{.j} = \frac{\sum_i \mu_{ij}}{I},$$

and

$$\gamma_{ij} = \mu_{ij} - (\mu + \alpha_i + \beta_j).$$

- It can be shown that defining the γ_{ij} 's, α_i 's, and β_j 's as on the previous slide is equivalent to imposing the constraints

$$\sum_i \alpha_i = 0 \quad \text{and} \quad \sum_j \beta_j = 0$$

and

$$\sum_j \gamma_{ij} = 0 \quad (\text{for each fixed } i) \quad \text{and} \quad \sum_i \gamma_{ij} = 0 \quad (\text{for each fixed } j).$$

Sums of Squares and the ANOVA Partition

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 - One reflecting variation **between** the levels of **Factor A**.
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 - The other reflecting variation **within** the groups.

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 - **SST** is the **total sum of squares**, defined as

$$\text{SST} = \sum_i \sum_j \sum_k (X_{ijk} - \bar{X} \dots)^2 \quad \text{df} = IJK - 1$$

which measures the **total** variation in the X_{ij} 's.

- (cont'd):
 - **SSA** is the **Factor A sum of squares**, defined as

$$\begin{aligned} \text{SSA} &= \sum_i \sum_j \sum_k (\bar{X}_{i..} - \bar{X}_{...})^2 \\ &= JK \sum_{i=1}^I (\bar{X}_{i.} - \bar{X}_{..})^2 \quad \mathit{df} = I - 1 \end{aligned}$$

which measures variation between the **levels of Factor A** due to both the **Factor A effect** and **random error**.

- (cont'd):
 - **SSB** is the **Factor B sum of squares**, defined as

$$\begin{aligned}
 \text{SSB} &= \sum_i \sum_j \sum_k (\bar{X}_{.j.} - \bar{X}_{...})^2 \\
 &= IK \sum_{i=1}^I (\bar{X}_{.j} - \bar{X}_{..})^2 \quad df = J - 1
 \end{aligned}$$

which measures variation between the **levels of Factor B** due to both the **Factor B effect** and **random error**.

- (cont'd):

- **SSAB** is the *interaction sum of squares*, with

$$\text{SSAB} = \sum_i \sum_j \sum_k (\bar{X}_{ij\cdot} - \bar{X}_{i\cdot\cdot} - \bar{X}_{\cdot j\cdot} + \bar{X}_{\dots})^2$$

$$df = (I - 1)(J - 1)$$

which measures variation due to an **interaction** (i.e. **non-additivity**) of the effects of the two factors and **random error**.

- (cont'd):
 - **SSE** is the **error sum of squares**, defined as

$$\text{SSE} = \sum_i \sum_j \sum_k (X_{ijk} - \bar{X}_{ij\cdot})^2$$

$$df = IJ(K - 1)$$

which measures variation of the X_{ijk} 's **within** treatment groups due to **random error**.

Proposition

ANOVA Partition for the Full Model: It can be shown that

$$SST = SSA + SSB + SSAB + SSE$$

Additive Property of Degrees of Freedom:

$$\text{df for SST} = \text{df for SSA} + \text{df for SSB} + \text{df for SSAB} + \text{df for SSE}$$

Mean Squares

- The Factor A mean square, Factor B mean square, interaction mean square, and mean squared error are:

$$MSA = \frac{SSA}{I - 1}$$

$$MSB = \frac{SSB}{J - 1}$$

$$MSAB = \frac{SSAB}{(I - 1)(J - 1)}$$

$$MSE = \frac{SSE}{IJ(K - 1)}$$

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 - Hypotheses about an ***interaction effect*** between the two factors:

$$H_{0AB} : \quad \gamma_{ij} = 0 \text{ for all } i \text{ and } j \quad (2)$$

$$H_{aAB} : \quad \text{Not all } \gamma_{ij} \text{'s equal zero}$$

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- Hypotheses about a ***Factor A main effect***:

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- Hypotheses about a **factor B main effect**:

$$H_{0B} : \quad \beta_j = 0 \text{ for all } j \quad (4)$$

$$H_{aB} : \quad \text{Not all } \beta_j \text{'s equal zero}$$

- In each case, the **null hypothesis** says there's **no effect** and the **alternative** says there **is an effect**.

- We always **test for an interaction effect first**, and proceed to the tests for main effects **only if the interaction effect isn't statistically significant**. We'll see why later.

Two-Factor ANOVA F Test Statistics:

$$F_{AB} = \frac{MSAB}{MSE} \quad F_A = \frac{MSA}{MSE} \quad F_B = \frac{MSB}{MSE}$$

Large values of F_{AB} provide evidence against H_{0AB} in favor of H_{aAB} : Not all of the γ_{ij} 's are zero.

Large values of F_A provide evidence against H_{0A} in favor of H_{aA} : Not all of the α_i 's are zero.

Large values of F_B provide evidence against H_{0B} in favor of H_{aB} : Not all of the β_j 's are zero.

- Suppose data in a two-factor study follow the **two-factor ANOVA model**, where the error terms ϵ_{ijk} are iid $N(0, \sigma)$.

Sampling Distributions of the Test Statistics Under H_0 :

1. If F_{AB} is the F test statistic for the interaction effect, then when

$$H_{0AB} : \gamma_{ij} = 0 \text{ for all } i \text{ and } j$$

is true,

$$F_{AB} \sim F((I - 1)(J - 1), IJ(K - 1)).$$

2. If F_A is the F test statistic for Factor A, then when

$$H_{0A} : \alpha_i = 0 \text{ for all } i$$

is true,

$$F \sim F(I - 1, IJ(K - 1)).$$

3. If F_B is the the F test statistic for Factor B, then when

$$H_{0B} : \beta_j = 0 \text{ for all } j$$

is true,

$$F \sim F(J - 1, IJ(K - 1)).$$

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 - The **rejection regions** as the **extreme largest $100\alpha\%$ of F values**.
 - The **p -values** as the **tail areas to the right of the observed F_{AB} , F_A , and F_B values**.

- **Comment:** The **ANOVA F tests** can be used even if the samples are from **non-normal** populations as long the per-group sample sizes are large.

The ANOVA Table

- The results are summarized in an **ANOVA table**:

Source of Variation	df	Sum of Squares	Mean Square	f	P-value
Factor A	$I - 1$	SSA	$MSA = SSA / (I - 1)$	$F_A = MSA / MSE$	p
Factor B	$J - 1$	SSB	$MSB = SSB / (J - 1)$	$F_B = MSB / MSE$	p
Interaction	$(I - 1)(J - 1)$	SSAB	$MSAB = SSAB / (I - 1)(J - 1)$	$F_{AB} = MSAB / MSE$	p
Error	$IJ(K - 1)$	SSE	$MSE = SSE / IJ(K - 1)$		
Total	$IJK - 1$	SST			

Example

For the study of factors effecting errors in predicting the completion time for a programming project, the **ANOVA table** is:

Source	df	Sum of Squares	Mean Square	f	P-value
Type	1	39447.0	39447.0	458.0	0.000
Years	2	36412.0	18206.0	211.4	0.000
Interaction	2	20165.3	10082.7	117.1	0.000
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(We'll interpret the results in a later example.)

Interpretation of the Interaction Effects γ_{ij}

- If the interaction effect $\gamma_{ij} = 0$ for all i and j , then the two-factor ANOVA model reduces to the so-called **additive effects** model

$$X_{ijk} = \mu + \alpha_i + \beta_j + \epsilon_{ijk}$$

- Under the ***additive effects*** model, the **effect of each factor**, at any **fixed level** of the **other factor**, is the **same regardless** of the **level** of that **other factor**.

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- By **including** the **interaction** term γ_{ij} in the model (and allowing it to be non-zero), we **allow** the **effect of each factor** to be **different** depending on the **level** of the **other factor**.

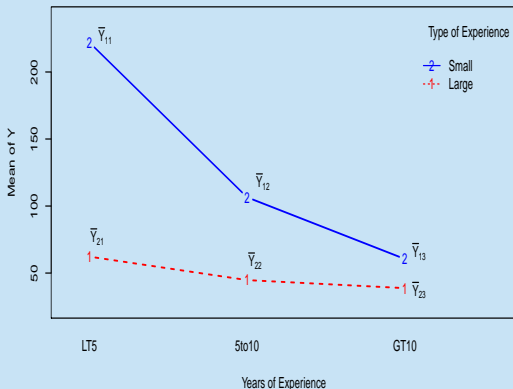
Interaction Plots

- An *interaction plot* is a plot of the **group means** \bar{X}_{ij} . on the *y*-axis, **levels of one factor** on the *x*-axis, and **lines** connecting **group means** at each **fixed level** of the **other factor**.

Example

For the data from the study of the effects of **type** of experience (Factor A) and **years** of experience (Factor B) on prediction errors, the **interaction plot** is below.

Interaction Plot for Programming Project Data



There appears to be an **interaction effect** between **type** and **years** of experience because the **effect** of **years** is **different** depending on the **type** of experience.

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Thus:

- **Lines that are close to parallel suggest** that the effect of Factor A is the same regardless of the level of Factor B, i.e. that **the effects of the two factors are *additive***.

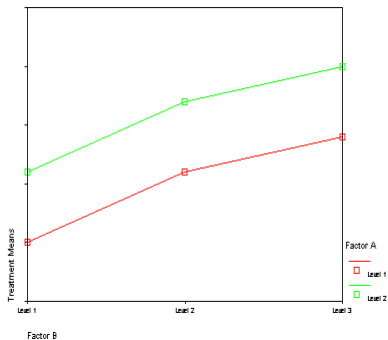
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Thus:

- **Lines that are close to parallel suggest** that the effect of Factor A is the same regardless of the level of Factor B, i.e. that **the effects of the two factors are *additive***.
- **Lines that aren't parallel suggest that there's an *interaction* effect between the two factors.**

Interaction Plot of Additive Model



Interaction Plot of Full Factorial Model

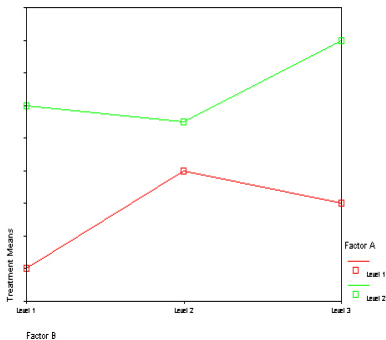


Figure: No interaction between two factors (i.e. their effects are *additive*) (left). Interactions between the two factors (i.e. the effects are *not additive*) (center and right).

Only Test for Main Effects if the Interaction Isn't Significant

- If we reject H_{0AB} , this **automatically** tells us that **both factors have effects, regardless** of whether or not we reject H_{0A} and H_{0B} .

Only Test for Main Effects if the Interaction Isn't Significant

- If we reject H_{0AB} , this **automatically** tells us that **both factors have effects, regardless** of whether or not we reject H_{0A} and H_{0B} .

But the **effect** of **each factor** is **different** depending on the **level** of the **other factor**.

Example

Here's the **ANOVA table** (again) for our running example:

Source	df	Sum of Squares	Mean Square	f	P-value
Type	1	39447.0	39447.0	458.0	0.000
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- There is an **interaction effect** between **type** of and **years** of experience (**p-value = 0.000**).
- There's **no need to proceed** to the tests for **main effects** because we know, since there's an **interaction**, that **regardless** of what their **p-values** are, both **type** and **years** of experience have effects.

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- There's **no need to proceed** to the tests for **main effects** because we know, since there's an **interaction**, that **regardless** of what their **p-values** are, both **type** and **years** of experience have effects.

Their effects each differ depending on the level of the other factor.

We can conclude:

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Their effects each differ depending on the level of the other factor.

The **interaction plot** depicts the **nature** of the **interaction effect**.

Main Effects *Masked* by an Interaction Effect

- We **only proceed** to tests for **main effects** if the interaction *isn't* significant.

Main Effects *Masked* by an Interaction Effect

- We **only proceed** to tests for **main effects** if the **interaction *isn't* significant**.

When the **interaction *is* significant**, the tests for main effects have **little practical meaning**.

Example

Consider the following data from a two-factor experiment.

		Factor B		
		Level 1	Level 2	
Factor A	Level 1	20 25	10 14	$\bar{X}_{1..} = 17.25$
	Level 2	11 14	21 24	$\bar{X}_{2..} = 17.50$
		$\bar{X}_{.1.} = 17.50$	$\bar{X}_{.2.} = 17.25$	

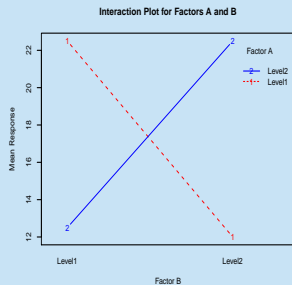
The resulting **ANOVA table** is below.

Source	df	Sum of Squares	Mean Square	f	P-value
Factor A	1	0.12	0.12	0.017	0.9027
Factor B	1	0.13	0.13	0.017	0.9027
Factor A:Factor B	1	210.13	210.13	28.492	0.0059
Error	4	29.50	7.37		
Total	7	239.88			

The **interaction** is **significant**, but **neither** of the **main effects** is.

The table and interaction plot below show that the Factor *A* and *B* main effects are **masked** by the interaction effect.

		Factor B		
		Level 1	Level 2	
Factor A	Level 1	$\bar{X}_{11.} = 22.5$	$\bar{X}_{12.} = 12.0$	$\bar{X}_{1..} = 17.25$
	Level 2	$\bar{X}_{21.} = 12.5$	$\bar{X}_{22.} = 22.5$	$\bar{X}_{2..} = 17.50$
		$\bar{X}_{.1.} = 17.50$	$\bar{X}_{.2.} = 17.25$	



- When the **interaction *is* significant**, we can investigate the effect of each factor **separately** for each fixed **level** of the **other factor**.

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One way to do this is to carry out **Tukey's multiple comparison procedure** on the IJ group means.