8 Two-Sample Hypothesis Tests

MTH 3240 Environmental Statistics

Spring 2020

イロト イポト イヨト イヨト

= 990

Objectives

Objectives:

- Carry out a two-sample *t* test for the difference between two population means.
- Compute and interpret two-sample *t* confidence interval for the difference between two population means.

イロン イボン イヨン イヨン

ъ

Introduction

• We're often interested in testing for a difference between **two** population means μ_x and μ_y .

イロト イポト イヨト イヨト

■ のへで

Introduction

• We're often interested in testing for a difference between **two** population means μ_x and μ_y .

イロト イポト イヨト イヨト

3

• Here are three examples.

In a **control-impact** study, we might suspect that the mean contaminant level is higher at the **impact site** than at the **control site**. The hypotheses would be

 $H_0: \mu_x = \mu_y$ $H_a: \mu_x > \mu_y$

where μ_x and μ_y are the true (unknown) **population mean** contaminant levels at the **impact** and **control sites**, respectively.

In a **control-impact** study, we might suspect that the mean contaminant level is higher at the **impact site** than at the **control site**. The hypotheses would be

 $H_0: \mu_x = \mu_y$ $H_a: \mu_x > \mu_y$

where μ_x and μ_y are the true (unknown) **population mean** contaminant levels at the **impact** and **control sites**, respectively.

 H_0 says there's **no difference** between the two sites' means, and H_a says the **impact site's** mean is **higher**.

In a **before-after** study, we might suspect that the site became contaminated as a result of the impact event. The hypotheses would be

$$H_0: \mu_x = \mu_y$$
$$H_a: \mu_x > \mu_y$$

< 口 > < 同 > < 臣 > < 臣 >

æ

where μ_x and μ_y are the true (unknown) mean contaminant levels **after** and **before** the impact event, respectively.

In a **before-after** study, we might suspect that the site became contaminated as a result of the impact event. The hypotheses would be

$$H_0: \mu_x = \mu_y$$
$$H_a: \mu_x > \mu_y$$

where μ_x and μ_y are the true (unknown) mean contaminant levels **after** and **before** the impact event, respectively.

 H_0 says the impact event had **no effect** on the contaminant levels, and H_a says it **increased** them.

イロト イポト イヨト イヨト

In an **experiment**, we might randomly assign experimental units to **treatment** and **control** conditions and compare their responses. To decide if there's **any difference** in the effects of the two conditions, we'd test

$$H_0: \mu_x = \mu_y$$
$$H_a: \mu_x \neq \mu_y$$

where μ_x and μ_y are the true (unknown) **population mean** responses to the two conditions.

In an **experiment**, we might randomly assign experimental units to **treatment** and **control** conditions and compare their responses. To decide if there's **any difference** in the effects of the two conditions, we'd test

 $H_0: \mu_x = \mu_y$ $H_a: \mu_x \neq \mu_y$

where μ_x and μ_y are the true (unknown) **population mean** responses to the two conditions.

 H_0 says there's **no difference** in the mean responses for the two conditions, and H_a says there's **a difference**.

• We'll look at two tests for comparing two population means:

イロト イポト イヨト イヨト

= 990

- 1. The two-sample t test
- 2. The rank sum test

- We'll look at two tests for comparing two population means:
 - 1. The two-sample t test
 - 2. The rank sum test

The t test requires a **normality** assumption (or large sample sizes), but the **rank sum test** is a **nonparametric** test (doesn't require normality).

ヘロト 人間 ト くほ ト くほ トー

æ

Two-Sample t Test

• For the *two-sample* t *test*, we suppose we have random samples

$$X_1, X_2, \dots, X_{n_x}$$
 and Y_1, Y_2, \dots, Y_{n_y}

from two populations whose means are

$$\mu_x$$
 and μ_y

イロト イポト イヨト イヨト

∃ 𝒫𝔅

Two-Sample t Test

• For the *two-sample* t *test*, we suppose we have random samples

$$X_1, X_2, \ldots, X_{n_x}$$
 and $Y_1, Y_2, \ldots, Y_{n_y}$

from two populations whose means are



• The sample sizes

 n_x and n_y

イロト イポト イヨト イヨト

ъ

don't have to be the same.

• The **null hypothesis** is that there's **no difference** between μ_x and μ_y .

Null Hypothesis:

$$H_0: \mu_x = \mu_y.$$

イロト 不同 とくほ とくほ とう

= 990

• The alternative hypothesis is one of the following.

Alternative Hypothesis:

- 1. $H_a: \mu_x > \mu_y$ (upper-tailed test)
- 2. $H_a: \mu_x < \mu_y$ (lower-tailed test)
- 3. $H_a: \mu_x \neq \mu_y$ (two-tailed test)

depending on what we're trying to verify using the data.

イロト イポト イヨト イヨト

æ

We'll usually write these hypotheses as

$$H_0: \mu_x - \mu_y = 0$$

and

1. $H_a: \mu_x - \mu_y > 0$ 2. $H_a: \mu_x - \mu_y < 0$ 3. $H_a: \mu_x - \mu_y \neq 0$ (upper-tailed test) (lower-tailed test) (two-tailed test)

イロト イポト イヨト イヨト

= 990

We'll usually write these hypotheses as

$$H_0: \mu_x - \mu_y = 0$$

and

- 1. $H_a: \mu_x \mu_y > 0$ (upper-tailed test)2. $H_a: \mu_x \mu_y < 0$ (lower-tailed test)3. $H_a: \mu_x \mu_y \neq 0$ (two-tailed test)
- The difference $\mu_x \mu_y$ between the population means is sometimes called the *effect size*.

イロト イポト イヨト イヨト

э.

We'll usually write these hypotheses as

$$H_0: \mu_x - \mu_y = 0$$

and

- 1. $H_a: \mu_x \mu_y > 0$ (upper-tailed test)2. $H_a: \mu_x \mu_y < 0$ (lower-tailed test)3. $H_a: \mu_x \mu_y \neq 0$ (two-tailed test)
- The difference $\mu_x \mu_y$ between the population means is sometimes called the *effect size*.

If μ_x and μ_y are true (unknown) mean responses to **treatment** and **control** conditions in an experiment, H_0 says the treatment has **no effect**.

ヘロト 人間 とくほとくほとう

- We'll denote the **sample means** by
 - $ar{X}$ and $ar{Y}$

and the sample standard deviations by

$$S_x$$
 and S_y

ヘロト 人間 とくほとくほとう

■ のへの

- We'll denote the **sample means** by
 - $ar{X}$ and $ar{Y}$

and the sample standard deviations by

 S_x and S_y

• We estimate the effect size by the difference between the sample means, $\bar{X} - \bar{Y}$.

イロン イボン イヨン イヨン



◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○



• The denominator of t, $S_{\bar{x}-\bar{y}}$, is an estimate of the **standard** error of the statistic $\bar{X} - \bar{Y}$.

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ● ●



- The denominator of t, $S_{\bar{x}-\bar{y}}$, is an estimate of the **standard** error of the statistic $\bar{X} \bar{Y}$.
- So t indicates how many standard errors the estimated effect size $\bar{X} \bar{Y}$ is away from zero.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

• $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...

- $ar{X} ar{Y}$ is an estimate of $\mu_x \mu_y$, so ...
 - If H_0 was true, ...

- $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...
 - If H_0 was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to be close to **zero**.

▲ロト ▲帰 ト ▲ 臣 ト ▲ 臣 ト ○ 臣 - の Q ()

- $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...
 - If H₀ was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to be close to **zero**.

▲ロト ▲帰 ト ▲ 臣 ト ▲ 臣 ト ○ 臣 - の Q ()

• But if Ha was true, ...

• $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...

• If H₀ was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to be close to **zero**.

• But if Ha was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to differ from zero in the direction specified by H_a .

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

- $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...
 - If H₀ was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to be close to **zero**.

• But if Ha was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to differ from zero in the direction specified by H_a .

▲□▶ ▲□▶ ▲目▶ ▲目▶ 目 のへで

• Thus ...

- $ar{X}-ar{Y}$ is an estimate of $\mu_x-\mu_y$, so ...
 - If H₀ was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to be close to **zero**.

• But if Ha was true, ...

... we'd expect $\bar{X} - \bar{Y}$ to differ from zero in the direction specified by H_a .

- Thus ...
 - 1. t will be approximately **zero** (most likely) if H_0 is true.

◆□ ▶ ◆□ ▶ ◆ □ ▶ ◆ □ ▶ ◆ □ ● ● ○ ○ ○

2. It will **differ from zero** (most likely) in the direction specified by H_a if H_a is true.

- 1. Large positive values of t provide evidence in favor of $H_a: \mu_x \mu_y > 0.$
- 2. Large negative values of t provide evidence in favor of $H_a: \mu_x \mu_y < 0.$
- 3. Both large positive and large negative values of t provide evidence in favor of $H_a: \mu_x \mu_y \neq 0$.

イロト イポト イヨト イヨト 一臣

Now suppose we have random samples from two populations.

Now suppose we have random samples from two populations.

If either

- The populations are both normal, or
- 2 The sample sizes n_x and n_y are both large,

イロト イポト イヨト イヨト

= 990

Now suppose we have random samples from two populations.

If either

- The populations are both normal, or
- 2 The sample sizes n_x and n_y are both large,

イロン イボン イヨン イヨン

3

the null distribution is as follows.

Sampling Distribution of *t* **Under** H_0 : If *t* is the two-sample *t* test statistic, then when

$$H_0: \mu_x - \mu_y = 0$$

is true,

 $t \sim t(\mathrm{df}),$

where the df are

$$\mathsf{df} \; = \; \frac{\left(\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}\right)^2}{\frac{(S_x^2/n_x)^2}{n_x - 1} + \frac{(S_y^2/n_y)^2}{n_y - 1}},$$

ヘロト 人間 とくほとくほとう

3

(which is rounded *down* to the nearest integer).

• **P-values** and **rejection regions** are obtained from the appropriate tail(s) of the t(df) distribution, as shown on the next slides.

イロト イポト イヨト イヨト

∃ <2 <</p>







Two-Sample t Test for μ_x and μ_y

Assumptions: The data $x_1, x_2, \ldots, x_{n_x}$ and $y_1, y_2, \ldots, y_{n_y}$ are independent random samples from two populations and either the populations are normal or n_x and n_y are large.

Null hypothesis: $H_0: \mu_x - \mu_y = 0.$

Test statistic value: $t = \frac{\bar{x} - \bar{y}}{\sqrt{s_x^2/n_x + s_y^2/n_y}}.$

Decision rule: Reject H_0 if p-value $< \alpha$ or t is in rejection region.

Two-Sample t Test for μ_x and μ_y				
		P-value = area under		
	Alternative	t distribution with	Rejection region =	
	hypothesis	d.f. given by df:	t values such that:	
	$H_a: \mu_x - \mu_y > 0$	to the right of t	$t > t_{lpha, df}$	
	$H_a: \mu_x - \mu_y < 0$	to the left of t	$t < -t_{lpha,df}$	
	$H_a: \mu_x - \mu_y \neq 0$	to the left of $- t $ and right of $ t $	$t > t_{lpha/2,df}$ or $t < -$	

* $t_{\alpha,\text{df}}$ is the $100(1-\alpha)$ th percentile of the *t* distribution with d.f. given by df.

Exercise

To assess the impact of a wastewater treatment plant's effluent discharge into the Febros River, Portugal, gudgeon fish (*Gobio gobio*) were sampled **upstream** and **downstream** of the plant and their **weights** (g) measured.

イロト イポト イヨト イヨト

э

Exercise

To assess the impact of a wastewater treatment plant's effluent discharge into the Febros River, Portugal, gudgeon fish (*Gobio gobio*) were sampled **upstream** and **downstream** of the plant and their **weights** (g) measured.

The water was warmer downstream than upstream, possibly due to the effluent discharge, and fish are more active and have higher metabolic activity in warmer water, which can lead to growth.

< □ > < 同 > < 三 > <

.⊒...>

Exercise

To assess the impact of a wastewater treatment plant's effluent discharge into the Febros River, Portugal, gudgeon fish (*Gobio gobio*) were sampled **upstream** and **downstream** of the plant and their **weights** (g) measured.

The water was warmer downstream than upstream, possibly due to the effluent discharge, and fish are more active and have higher metabolic activity in warmer water, which can lead to growth.

Also, the nutrient load was higher downstream, and this can lead to more available food for the fish.

イロト イポト イヨト イヨト

イロト 不得 とくほと くほとう

イロト 不得 とくほ とくほとう

ъ

The summary statistics are below.

The summary statistics are below.

Gudgeon Weights		
Upstream	Downstream	
$n_x~=~23$	$n_y = 22$	
$ar{X}~=~12.12$	$ar{Y}~=~16.93$	
$S_x\ =\ 2.57$	$S_y = 4.17$	

イロト 不得 とくほ とくほとう

ъ

The summary statistics are below.

Gudgeon Weights			
Upstream	Downstream		
$n_x~=~23$	$n_y = 22$		
$ar{X}~=~12.12$	$ar{Y}~=~16.93$		
$S_x\ =\ 2.57$	$S_y = 4.17$		

Carry out a **two-sample** *t* **test** to decide if the population mean gudgeon weight is **greater downstream than upstream**. Use a level of significance $\alpha = 0.05$.

イロト イポト イヨト イヨト

Hints: The degrees of freedom are

df =
$$\frac{\left(\frac{2.57^2}{23} + \frac{4.17^2}{22}\right)^2}{\frac{(2.57^2/23)^2}{23-1} + \frac{(4.17^2/22)^2}{22-1}} = 34,$$

▲□▶▲□▶▲□▶▲□▶ □ のQの

and you should get t = -4.63 and **p-value = 0.0000**.

Two-Sample t Confidence Interval

• We can compute a *two-sample* t *confidence interval* for the true (unknown) effect size $\mu_x - \mu_y$.

◆□▶ ◆□▶ ◆三▶ ◆三▶ ● ● ●

Two-Sample t Confidence Interval

 We can compute a *two-sample* t confidence interval for the true (unknown) effect size μ_x - μ_y.

Two-Sample t CI: A $100(1 - \alpha)\%$ two-sample t CI for $\mu_x - \mu_y$ is $\bar{X} - \bar{Y} \pm t_{\alpha/2,df} S_{\bar{X}-\bar{Y}}$ where $S_{\bar{X}-\bar{Y}} = \sqrt{\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}},$

イロト イポト イヨト イヨト

and the **df** are the same as for the two-sample *t* test.

The CI is valid if either

- The two populations are both normal, or
- 2 The sample sizes n_x and n_y are both large.

イロト イポト イヨト イヨト

∃ <2 <</p>

The CI is valid if either

The two populations are both normal, or

2 The sample sizes n_x and n_y are both large.

• We can be $100(1 - \alpha)\%$ confident that the true (unknown) effect size $\mu_x - \mu_y$ will be contained in the interval.

イロト イポト イヨト イヨト

= 990

• The "plus or minus" part is called the *margin of error*.

• The "plus or minus" part is called the *margin of error*.



イロト イポト イヨト イヨト

• The "plus or minus" part is called the *margin of error*.



A smaller margin of error indicates that X
 - Y
 is a more precise estimate of the (unknown) effect size μ_x – μ_y.

・ロン・西方・ ・ ヨン・ ヨン・

For the study of impact of the wastewater treatment plant's discharge into the Febros River on fish weights, recall that the summary statistics are:

Gudgeon Weights			
Upstream	Downstream		
$n_x = 23$	$n_y~=~22$		
$ar{X}~=~12.12$	$ar{Y}~=~16.93$		
$S_x\ =\ 2.57$	$S_y \ = \ 4.17$		

イロト イポト イヨト イヨト

The estimate of the true (unknown) effect size $\mu_x - \mu_y$ is $\bar{X} - \bar{Y} = 12.12 - 16.93 = -4.81$,

The estimate of the true (unknown) effect size $\mu_x - \mu_y$ is

$$\bar{X} - \bar{Y} = 12.12 - 16.93 = -4.81,$$

ヘロト 人間 とくほとく ほとう

= 990

i.e. we estimate that gudgeon weights are **4.81 g** heavier **downstream** than **upstream**, on average.

500

イロ とうゆ とう ちょう ちょう

$$S_{\bar{x}-\bar{y}} = \sqrt{\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}} = \sqrt{\frac{2.57^2}{23} + \frac{4.17^2}{22}} = 1.04.$$

500

ヘロン ヘロ・シャー ショー・

$$S_{\bar{x}-\bar{y}} = \sqrt{\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}} = \sqrt{\frac{2.57^2}{23} + \frac{4.17^2}{22}} = 1.04.$$

so 95% two-sample t CI for $\mu_x - \mu_y$ is

$$S_{\bar{x}-\bar{y}} = \sqrt{\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}} = \sqrt{\frac{2.57^2}{23} + \frac{4.17^2}{22}} = 1.04.$$

so 95% two-sample t CI for $\mu_x - \mu_y$ is

$$\begin{split} \bar{X} - \bar{Y} &\pm t_{\alpha/2, \text{df}} S_{\bar{X} - \bar{Y}} &= -4.81 \pm 2.032 \, (1.04) \\ &= -4.81 \pm 2.11 \\ &= (-6.92, -2.70) \end{split}$$

$$S_{\bar{x}-\bar{y}} = \sqrt{\frac{S_x^2}{n_x} + \frac{S_y^2}{n_y}} = \sqrt{\frac{2.57^2}{23} + \frac{4.17^2}{22}} = 1.04.$$

so 95% two-sample t CI for $\mu_x - \mu_y$ is

$$\bar{X} - \bar{Y} \pm t_{\alpha/2, \text{df}} S_{\bar{X} - \bar{Y}} = -4.81 \pm 2.032 (1.04)$$
$$= -4.81 \pm 2.11$$
$$= (-6.92, -2.70)$$

(where the *t* critical value $t_{0.025,df} = 2.032$ was obtained from a *t* distribution table and the df is 34 from the previous exercise).

We're **95% confident** that the true (unknown) effect of the effluent discharge on gudgeon weights is an **increase** of between **2.70** and **6.92 g**.

イロト イポト イヨト イヨト

We're **95% confident** that the true (unknown) effect of the effluent discharge on gudgeon weights is an **increase** of between **2.70** and **6.92 g**.

From the previous slide, the **margin of error** in the **estimate** of the effect size is **2.11 g**.

イロト イポト イヨト イヨト