

11 Tests for the Effects of Two Factors (Cont'd)

MTH 3240 Environmental Statistics

Spring 2020

Objectives

Objectives:

- Write out the group means and treatment effects versions of the two-factor ANOVA model, including any assumptions about the random error term ϵ . (**Optional for Spring 2020**)
- Interpret sums of squares, degrees of freedom, and mean squares in two-factor ANOVA.
- Carry out two-factor ANOVA F tests for the effects of two factors and their interaction effect.
- Interpret the interaction effect in a two-factor study.
- Obtain and interpret fitted values and residuals in two-factor ANOVA.
- Use plots to check the normality and common population standard deviation assumptions required by the ANOVA F tests.

Two-Factor ANOVA Models (Optional for Spring 2020)

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The population standard deviations are required to all the same, σ .

(Optional for Spring 2020)

		Factor B				
		Level $j = 1$	Level $j = 2$...	Level $j = b$	
Factor A	Level $i = 1$	$\left. \begin{array}{c} Y_{111} \\ Y_{112} \\ \vdots \\ Y_{11n} \end{array} \right\} \sim N(\mu_{11}, \sigma)$	$\left. \begin{array}{c} Y_{121} \\ Y_{122} \\ \vdots \\ Y_{12n} \end{array} \right\} \sim N(\mu_{12}, \sigma)$...	$\left. \begin{array}{c} Y_{1b1} \\ Y_{1b2} \\ \vdots \\ Y_{1bn} \end{array} \right\} \sim N(\mu_{1b}, \sigma)$	$\mu_{1\cdot}$
	Level $i = 2$	$\left. \begin{array}{c} Y_{211} \\ Y_{212} \\ \vdots \\ Y_{21n} \end{array} \right\} \sim N(\mu_{21}, \sigma)$	$\left. \begin{array}{c} Y_{221} \\ Y_{222} \\ \vdots \\ Y_{22n} \end{array} \right\} \sim N(\mu_{22}, \sigma)$...	$\left. \begin{array}{c} Y_{2b1} \\ Y_{2b2} \\ \vdots \\ Y_{2bn} \end{array} \right\} \sim N(\mu_{2b}, \sigma)$	$\mu_{2\cdot}$

	Level $i = a$	$\left. \begin{array}{c} Y_{a11} \\ Y_{a12} \\ \vdots \\ Y_{a1n} \end{array} \right\} \sim N(\mu_{a1}, \sigma)$	$\left. \begin{array}{c} Y_{a21} \\ Y_{a22} \\ \vdots \\ Y_{a2n} \end{array} \right\} \sim N(\mu_{a2}, \sigma)$...	$\left. \begin{array}{c} Y_{ab1} \\ Y_{ab2} \\ \vdots \\ Y_{abn} \end{array} \right\} \sim N(\mu_{ab}, \sigma)$	$\mu_{a\cdot}$
		$\mu_{\cdot 1}$	$\mu_{\cdot 2}$...	$\mu_{\cdot b}$	μ

(Optional for Spring 2020)

- The value of μ_{ij} will depend on the level i of factor A and the level j of factor B .

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- We can describe the data using the **statistical model** called the ***additive effects two-factor ANOVA model***.

In the model, each μ_{ij} is written in terms of an **effect** of level i of factor A and an **effect** of level j of factor B .

Additive Two-Factor ANOVA Model (Optional for Spring 2020)

Additive Two-Factor ANOVA Model: One statistical model for describing data in a two-factor study is:

$$Y_{ijk} = \underbrace{\mu + \alpha_i + \beta_j}_{\text{This is } \mu_{ij}} + \epsilon_{ijk},$$

where

Y_{ijk} is the k th observation at the i th level of factor A and j th level of factor B .

μ is a constant called the **overall true mean**.

α_i is the **effect** of the i th level of factor A .

β_j is the **effect** of the j th level of factor B .

ϵ_{ijk} is a $N(0, \sigma)$ **random error** term.

(Optional for Spring 2020)

Example

For the soil phosphorus study, the **additive effects model** for describing a phosphorus concentration Y is of the form

$$Y = \text{Overall Mean} + \text{Soil Type Effect} + \text{Topography Effect} + \text{Error}$$

- But the **additive effects model** requires that the **effects** of the two factors be *additive*.

(Optional for Spring 2020)

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- But the **additive effects model** requires that the **effects** of the two factors be **additive**.
- They're **additive** when the effect of **factor A** is the **same regardless** of the **level** of **factor B**, and the effect of factor **B** is the **same regardless** of the **level** of **factor A**.

Interactions

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Interactions

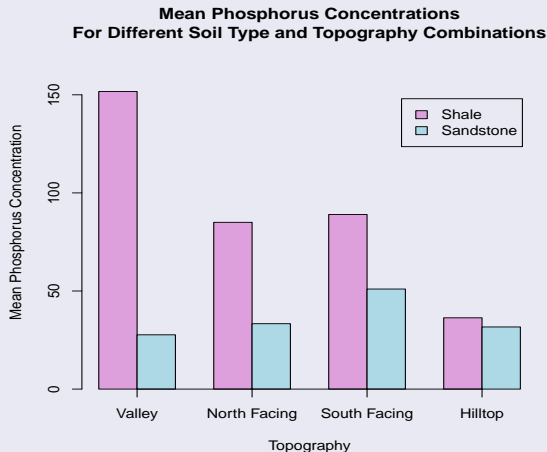
- When the effects of two factors **aren't** additive, we say that there's an ***interaction effect*** between them.
- An ***interaction effect*** exists when the effect of **factor A** is **different** depending on the **level** of **factor B** , and the effect of **factor B** is **different** depending on the **level** of **factor A** .
- We can check for an **interaction effect** by graphing the **group means** either in a **bar plot** or in an ***interaction plot***.

Example

For the soil phosphorus study, the eight **group means** are shown in the **body** of the table below.

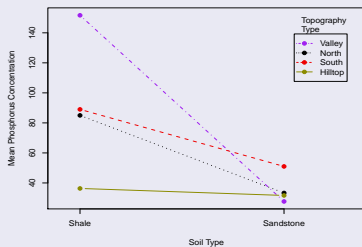
		Factor B: Topography				
		Valley (j=1)	North- Facing (j=2)	South- Facing (j=3)	Hilltop (j=4)	
Factor A: Soil Type	Shale (i=1)	$\bar{Y}_{11} =$ 151.7	$\bar{Y}_{12} =$ 85.0	$\bar{Y}_{13} =$ 89.0	$\bar{Y}_{14} =$ 36.3	$\bar{Y}_{1.} =$ 90.5
	Sand- stone (i=2)	$\bar{Y}_{21} =$ 27.7	$\bar{Y}_{22} =$ 33.3	$\bar{Y}_{23} =$ 51.0	$\bar{Y}_{24} =$ 31.7	$\bar{Y}_{2.} =$ 35.9
		$\bar{Y}_{.1} =$ 89.7	$\bar{Y}_{.2} =$ 59.2	$\bar{Y}_{.3} =$ 70.0	$\bar{Y}_{.4} =$ 34.0	$\bar{Y} =$ 63.2

A **bar plot** of the **group means** is shown (again) below.

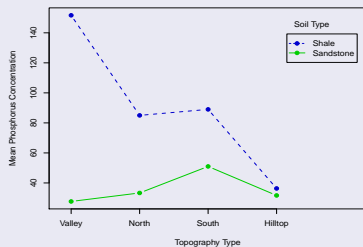


Here are **interaction plots** of the **group means**.

Interaction Plot of Soil Type and Topography Type

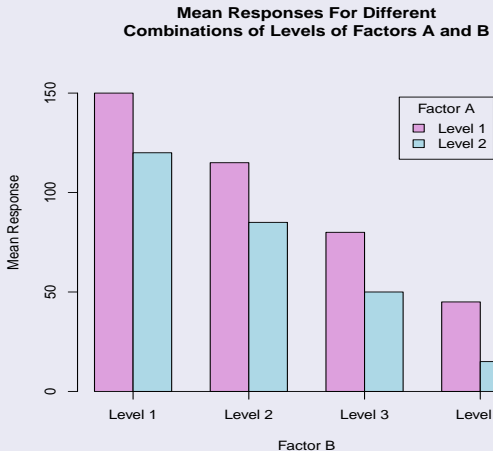


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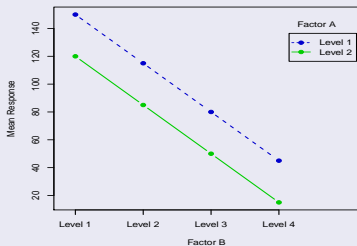
Example

Here's **bar plot** showing group means for two factors whose effects are **additive**.

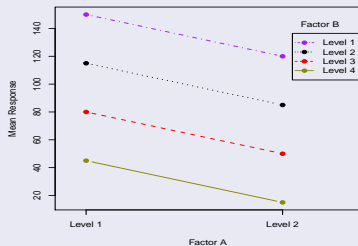


Here are **interaction plots** of those same **group means** (for two factors whose effects are **additive**).

Interaction Plot of Factors A and B



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 - We only have to look at one of the **interaction plots** or the other (not both).

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- In practice:
 - We only have to look at one of the **interaction plots** or the other (not both).
 - The lines in an **interaction plot** won't ever be **exactly** parallel because of random **sampling error**.

Two-Factor ANOVA Model With Interaction (Optional for Spring 2020)

- The model we'll use when the effects of two factors **aren't** additive is the ***two-factor ANOVA model with interaction effect***.

Two-Factor ANOVA Model With Interaction Effect: Another statistical model for describing data in a two-factor study is:

$$Y_{ijk} = \underbrace{\mu + \alpha_i + \beta_j + (\alpha\beta)_{ij}}_{\text{This is } \mu_{ij}} + \epsilon_{ijk},$$

(Optional for Spring 2020)

where

Y_{ijk} is the k th observation at the i th level of factor A
and j th level of factor B .

μ is a constant called the **overall true mean**.

α_i is the **main effect** of the i th level of factor A .

β_j is the **main effect** of the j th level of factor B .

$(\alpha\beta)_{ij}$ is an **interaction effect** representing **non-additivity** in the effects of the two factors.

ϵ_{ijk} is a $N(0, \sigma)$ **random error** term.

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... and it can still be used even if they ***are additive***.

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So it's "**safer**" to **include** the **interaction effect** in the **model** when carrying out a **two-factor ANOVA**.

From here on, we'll ***always* include** the **interaction effect** in the model.

Example

For the soil phosphorus study, the interaction plots suggest that the effects of **soil type** and **topography *aren't* additive**, so the we'd definitely use the **ANOVA model *with* the interaction effect**.

For a phosphorus concentration Y , the model has the form:

$$Y = \text{Overall Mean} + \text{Soil Type Effect} + \text{Topography Effect} \\ + \text{Interaction Effect} + \text{Error}$$

Fitted Values and Residuals

- The **group means** $\bar{Y}_{11}, \bar{Y}_{12}, \dots, \bar{Y}_{ab}$ are sometimes called ***fitted values***.

Fitted Values:

$$\text{Fitted Value for } ij\text{th Group} = \bar{Y}_{ij}$$

- Comments:**
 - Statistical software reports ***n* duplicates** of the **fitted value** for **each** of the ***ab* groups**, one duplicate for each of the ***n*** individuals in the group.

Residuals

- A **residual**, denoted e_{ijk} , is the **deviation** of an individual's observed Y_{ijk} value away from the **fitted value** for that individual.

Residuals:

$$e_{ijk} = Y_{ijk} - \bar{Y}_{ij}$$

- Statistical software reports the values of all N **residuals**, one for each individual in the study.

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(We'll also test for an **interaction effect** by comparing **variation due to the interaction** to **within-groups variation**.)

- We measure the different types of variation using ***sums of squares*** (and later ***mean squares***).

- We measure the **between-rows variation** by the **factor A sum of squares**, denoted **SSA**.

Factor A Sum of Squares:

$$SSA = nb \sum_{i=1}^a (\bar{Y}_{i\cdot} - \bar{Y})^2.$$

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Factor A Sum of Squares:

$$SSA = nb \sum_{i=1}^a (\bar{Y}_{i.} - \bar{Y})^2.$$

SSA will be **large** when there's substantial variation among **row means** $\bar{Y}_{1.}, \bar{Y}_{2.}, \dots, \bar{Y}_{a.}$, which would suggest **factor A** has an **effect**.

- We measure the **between-columns variation** by the ***factor B sum of squares***, denoted **SSB**.

Factor B Sum of Squares:

$$\text{SSB} = na \sum_{j=1}^b (\bar{Y}_{.j} - \bar{Y})^2.$$

- We measure the **between-columns variation** by the **factor B sum of squares**, denoted **SSB**.

Factor B Sum of Squares:

$$\text{SSB} = na \sum_{j=1}^b (\bar{Y}_{.j} - \bar{Y})^2.$$

SSB will be **large** when there's substantial variation among **column means** $\bar{Y}_{.1}, \bar{Y}_{.2}, \dots, \bar{Y}_{.b}$, which would suggest **factor B** has an **effect**.

- We measure **variation** due to an **interaction effect** by the *AB interaction sum of squares*, denoted **SSAB**.

***AB* Interaction Sum of Squares:**

$$\text{SSAB} = n \sum_{i=1}^a \sum_{j=1}^b (\bar{Y}_{ij} - \bar{Y}_{i\cdot} - \bar{Y}_{\cdot j} + \bar{Y})^2.$$

- We measure **variation** due to an **interaction effect** by the ***AB interaction sum of squares***, denoted **SSAB**.

AB Interaction Sum of Squares:

$$SSAB = n \sum_{i=1}^a \sum_{j=1}^b (\bar{Y}_{ij} - \bar{Y}_{i.} - \bar{Y}_{.j} + \bar{Y})^2.$$

It can be shown that SSAB will be **large** when the **group means** $\bar{Y}_{11}, \bar{Y}_{12}, \dots, \bar{Y}_{ab}$ aren't consistent with an additive effects model, which would suggest there's an **interaction effect** between **factors A** and **B** (i.e. their effects aren't additive).

- We measure **within-groups variation** by the ***error sum of squares***, denoted **SSE**.

Error Sum of Squares:

$$\text{SSE} = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^n (Y_{ijk} - \bar{Y}_{ij})^2 = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^n e_{ijk}^2.$$

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The **error sum of squares is just the *sum of squared residuals***, and reflects variation due to **random error**.

SSE will be **large** if there's **substantial** variation among individual observations (Y_{ijk} 's) **within** groups, and **small** otherwise.

The ANOVA Partition

- The **total sum of squares**, denoted **SSTo**, measures **total variation** in the data.

Total Sum of Squares:

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Total Sum of Squares:

$$SSTo = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^n (Y_{ijk} - \bar{Y})^2 .$$

SSTo reflects **total variation** due both to effects of the factors (if they have any effects) **and** random variation among individual observations within groups.

- It can be shown that

Two-Factor ANOVA Partition:

$$SSTo = SSA + SSB + SSAB + SSE.$$

This splits the total variation in the data as:

$$\begin{aligned} \text{Total Variation} &= \text{Between-Rows Variation} \\ &+ \text{Between-Columns Variation} \\ &+ \text{Variation Due to Interaction} \\ &+ \text{Within-Groups Variation} \end{aligned}$$

Example

For the soil phosphorus study, statistical software reports the following **sums of squares**.

$$SSTo = 51406.0 \quad (\text{Total variation})$$

$$SSA = 17876.0 \quad (\text{Variation due to soil type})$$

$$SSB = 9693.8 \quad (\text{Variation due to topography})$$

$$SSAB = 11390.8 \quad (\text{Variation due to interaction})$$

$$SSE = 12445.3 \quad (\text{Variation due to random error})$$

We see that the **two-factor ANOVA partition** holds since

$$51406.0 = 17876.0 + 9693.8 + 11390.8 + 12445.3.$$

This indicates that a large portion of the total variation in phosphorus measurements (**17876.0** out of **51406.0**, or **35%**) is due to the difference between the row means (soil types).

Degrees of Freedom

- Here are the **degrees of freedom** associated with the sums of squares in two-factor ANOVA. (These will determine which F distributions our p-values come from.)

Degrees of Freedom: For two-factor ANOVA, the degrees of freedom are:

$$df \text{ for SSTo} = N - 1$$

$$df \text{ for SSA} = a - 1$$

$$df \text{ for SSB} = b - 1$$

$$df \text{ for SSAB} = (a - 1)(b - 1)$$

$$df \text{ for SSE} = ab(n - 1) = N - ab$$

- The degrees of freedom, like the associated sums of squares, are additive in the following sense.

Additivity of Degrees of Freedom:
$$df \text{ for } SSTo = df \text{ for } SSA + df \text{ for } SSB + df \text{ for } SSAB + df \text{ for } SSE.$$

Example

For the soil phosphorus study, we have $a = 2$ soil types, $b = 4$ topographies, and $n = 3$ phosphorus observations per group. Thus the total number of phosphorus observations (overall sample size) is $N = 24$, and

$$df \text{ for SSTo} = 23$$

$$df \text{ for SSA} = 1$$

$$df \text{ for SSB} = 3$$

$$df \text{ for SSAB} = 3$$

$$df \text{ for SSE} = 16.$$

As expected,

$$23 = 1 + 3 + 3 + 16.$$

Mean Squares

- A **mean square** is a **sum of squares** divided by its **degrees of freedom**.
- The **factor A mean square**, **factor B mean square**, **AB interaction mean square**, and **mean squared error** are below.

Mean Squares: For two-factor ANOVA, the mean squares are

$$MSA = \frac{SSA}{a-1}$$

$$MSB = \frac{SSB}{b-1}$$

$$MSAB = \frac{SSAB}{(a-1)(b-1)}$$

$$MSE = \frac{SSE}{ab(n-1)}$$

The ANOVA F Tests

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In each case, H_0 says there's **no effect**, and H_a say's there's **an effect**.

1. Test for a **factor A main effect**:

H_{OA} : There's no factor A effect

H_{aA} : There is a factor A effect

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H_{OB} : There's no factor B effect

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H_{OB} : There's no factor B effect

H_{aB} : There is a factor B effect

3. Test for an **AB interaction effect**:

H_{OAB} : There's no factor A and B interaction effect

H_{aAB} : There is a factor A and B interaction effect

(Optional for Spring 2020)

In terms of the **ANOVA model parameters** these are written as:

1. Test for a ***factor A main effect***:

$$H_{OA} : \quad \alpha_1 = \alpha_2 = \cdots = \alpha_a = 0$$

$$H_{aA} : \quad \text{The } \alpha_i \text{'s don't all equal 0}$$

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2. Test for a **factor B main effect**:

$$H_{OB} : \quad \beta_1 = \beta_2 = \cdots = \beta_b = 0$$

$$H_{aB} : \quad \text{The } \beta_j \text{'s don't all equal 0}$$

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1. Test for a **factor A main effect**:

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2. Test for a **factor B main effect**:

$$H_{OB} : \quad \beta_1 = \beta_2 = \cdots = \beta_b = 0$$

$$H_{aB} : \quad \text{The } \beta_j\text{'s don't all equal 0}$$

3. Test for an **AB interaction effect**:

$$H_{OAB} : \quad (\alpha\beta)_{11} = (\alpha\beta)_{12} = \cdots = (\alpha\beta)_{ab} = 0$$

$$H_{aAB} : \quad \text{The } (\alpha\beta)_{ij}\text{'s don't all equal 0}$$

- Here are the corresponding ***two-factor ANOVA F test statistics***:

Two-Factor ANOVA F Test Statistics:

$$F_A = \frac{MSA}{MSE}$$

$$F_B = \frac{MSB}{MSE}$$

$$F_{AB} = \frac{MSAB}{MSE}$$

- Not that we can think of

$$F_A = \frac{\text{Between-Rows Variation}}{\text{Within-Groups Variation}}.$$

$$F_B = \frac{\text{Between-Columns Variation}}{\text{Within-Groups Variation}}.$$

$$F_{AB} = \frac{\text{Variation Due To Interaction}}{\text{Within-Groups Variation}}.$$

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In each case, *large* values of F (larger than about 1) provide evidence in favor of H_a .

- Now suppose the ab groups are samples from **normal** populations that all have the **same standard deviation σ** (or that they have the **same σ** and the common sample size n is **large**).

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In this case, the **null distributions** are as follows.

Sampling Distribution of F Under H_0 : If F_A , F_B , and F_{AB} are the two-factor ANOVA F test statistics, then:

- 1 When H_{OA} is true,

$$F_A = \frac{MSA}{MSE} \sim F(a - 1, N - ab).$$

- 2 When H_{OB} is true,

$$F_B = \frac{MSB}{MSE} \sim F(b - 1, N - ab).$$

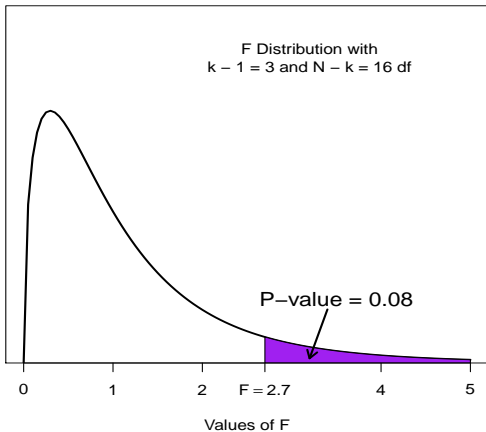
- 3 When H_{OAB} is true,

$$F_{AB} = \frac{MSAB}{MSE} \sim F((a - 1)(b - 1), N - ab).$$

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- The next slide shows the **p-value** when the observed test statistic value is $F = 2.7$.

P-Value for ANOVA F Test



The ANOVA Table

- The results of an **analysis of variance** are summarized in a ***two-factor ANOVA table*** having the form shown below.

Two-Factor ANOVA Table:

Source	DF	SS	MS	F	P-value
Factor A	$a - 1$	SSA	$MSA = SSA / (a - 1)$	MSA / MSE	p
Factor B	$b - 1$	SSB	$MSB = SSB / (b - 1)$	MSB / MSE	p
Interaction	$(a - 1)(b - 1)$	SSAB	$MSAB = SSAB / ((a - 1)(b - 1))$	$MSAB / MSE$	p
Error	$ab(n - 1)$	SSE	$MSE = SSE / (ab(n - 1))$		
Total	$N - 1$	SSTo			

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(But the effect of each factor is different depending on the level of the other factor.)
 - 2 If the AB interaction **isn't significant**, **proceed** to the results of the tests for **factor A and B main effects**.

Example

For the soil phosphorus study, the **ANOVA table** (obtained using software) is below.

Source of Variation	df	Sum of Squares	Mean Square	f	P-value
Soil Type	1	17876.0	17876.0	22.98	0.000
Topography Type	3	9693.8	3231.3	4.15	0.024
Interaction	3	11390.8	3796.9	4.88	0.013
Error	16	12445.3	777.8		
Total	23	51406.0			

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What **conclusions** are appropriate (using $\alpha = 0.05$)?

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We **already know** (because the interaction is significant) that each has an effect on phosphorus, **regardless** of what their p-values are.

Each factor has a **different effect** depending on the **level** of the **other factor**.

Checking the ANOVA Assumptions

- The **ANOVA F tests** require that the ab groups (samples) are from **normal** populations (or that their sample sizes are **large**) whose **standard deviations** are all **equal**.

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 - Make *ab separate histograms* or **normal probability plots**, one for each of the *ab groups*.
 - Make a *single histogram* or **normal probability plot** plot of the *N residuals* e_{ijk} .

- A few ways to check the **equal population standard deviation** assumption:

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 - An ***individual value plot*** of the ab groups.

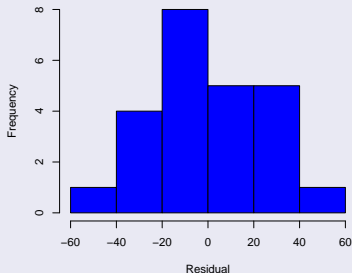
- A few ways to check the **equal population standard deviation** assumption:
 - An **individual value plot** of the ab groups.
 - A plot of the **residuals** (y -axis) versus **fitted values** (group means, x -axis).

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 - An **individual value plot** of the ab groups.
 - A plot of the **residuals** (y -axis) versus **fitted values** (group means, x -axis).
- In both plots, we look for roughly **equal amounts of within-group (vertical) spread** across the ab groups.

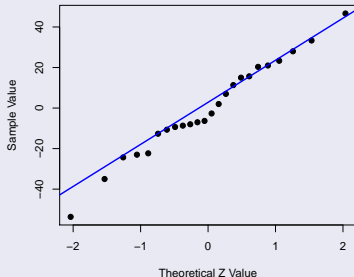
Example

For the soil phosphorus study, a **histogram** and **normal probability plot** of the **residuals** are below.

Histogram of Residuals



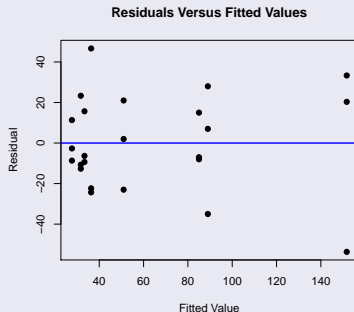
Normal Probability Plot of Residuals



The plots show that the **normality assumption** appears to be met.

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A plot the **residuals** versus the **fitted values (group means)** is below.



The amount of (vertical) spread of the points is roughly the same from left to right, so the **equal standard deviation assumption** appears to be met.

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Because the **normality** and **equal standard deviation assumptions** are met, the results of the F tests are **valid**.