

Statistical Methods

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Topics

1 Two-Factor ANOVA with $K = 1$ (Cont'd)

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Objectives

Objectives:

- Carry out Tukey's multiple comparison procedure after a two-factor ANOVA with $K = 1$, and interpret the results.
- Give the definition of a randomized block experiment, state the goal of randomized block experiments and describe their advantage over completely randomized experiments.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Two-Factor ANOVA with $K = 1$ (Cont'd)

Multiple Comparisons in the Additive Effects Model

- After rejecting either H_{0A} or H_{0B} , **Tukey's procedure** can be used to determine *which levels* of the factor **differ**.

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Tukey's Multiple Comparison Procedure: After the two-factor ANOVA F test rejects H_{0A} or H_{0B} :

1. Choose an **overall familywise confidence level** $100(1 - \alpha)\%$ (usually $\alpha = 0.05$ for a 95% confidence level).
2. For **Factor A comparisons**, compute the $I(I - 1)/2$ **CIs**:

$$\bar{X}_{i.} - \bar{X}_{j.} \pm Q_{\alpha, I, I-I-J+1} \sqrt{\frac{MSE}{J}}.$$

For **Factor B comparisons**, compute the $J(J - 1)/2$ **CIs**:

$$\bar{X}_{.j} - \bar{X}_{.j'} \pm Q_{\alpha, J, J-I-J+1} \sqrt{\frac{MSE}{I}}.$$

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Two-Factor ANOVA with $K = 1$ (Cont'd)

3. For any interval that **doesn't contain zero**, deem those **levels** of the given factor to be **different**.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Example

For the study of the effects of **brand** and **storage time** on vitamin C in orange juice, we found that **storage time** had an **effect**, but **brand** didn't.

The **Tukey procedure** in R produces the following CIs:

Times	Difference	Lower End Pt	Upper End Pt
Day3-Day7	0.63	-7.779	9.046
Day0-Day7	10.30	1.887	18.713 *
Day0-Day3	9.67	1.254	18.079 *

Intervals marked with asterisks don't contain zero.

We conclude that **Day 0** differs from both **Days 3** and **7**, but **Days 3** and **7** don't differ from each other.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Estimating Parameters in the Additive Effects Model

- Recall that the *additive effects version* of the two-factor ANOVA model is:

$$X_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}. \quad (1)$$

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Model Parameter Estimators: We estimate the unknown model parameters μ , α_i , β_j , and σ using the **estimators** $\hat{\mu}$, $\hat{\alpha}_i$, $\hat{\beta}_j$, and $\hat{\sigma}$ defined as:

Model Parameter	Estimator
μ	$\hat{\mu} = \bar{X}_{..}$
$\alpha_i = \mu_i - \mu$	$\hat{\alpha}_i = \bar{X}_i - \bar{X}_{..}$
$\beta_j = \mu_{.j} - \mu$	$\hat{\beta}_j = \bar{X}_{.j} - \bar{X}_{..}$
σ	$\hat{\sigma} = \sqrt{MSE}$

Predicted Values and Residuals for the Additive Effects Model

- The **fitted value** (or **predicted value**) for the individual in the i, j th cell, \hat{X}_{ij} , is defined as:

$$\begin{aligned}\hat{X}_{ij} &= \hat{\mu} + \hat{\alpha}_i + \hat{\beta}_j \\ &= \bar{X}_{..} + (\bar{X}_i - \bar{X}_{..}) + (\bar{X}_{.j} - \bar{X}_{..}) \\ &= \bar{X}_i + \bar{X}_{.j} - \bar{X}_{..}\end{aligned}$$

\hat{X}_{ij} is the value we'd predict, based on the data, for the response of the individual in the i, j th cell.

- The **residual** for the observation in the i, j th cell, e_{ij} , is defined as

$$\begin{aligned}e_{ij} &= X_{ij} - \hat{X}_{ij} \\ &= X_{ij} - (\hat{\mu} + \hat{\alpha}_i + \hat{\beta}_j) \\ &= X_{ij} - \bar{X}_i - \bar{X}_{.j} + \bar{X}_{..}\end{aligned}$$

The **residual** e_{ij} corresponds to the **random error** term ϵ_{ij} in the model.

- Comment:** The **error sum of squares** (Slides 13) is the **sum of squared residuals**, i.e.

$$SSE = \sum_i \sum_j e_{ij}^2.$$

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Randomized Block Experiments

- In a **one-factor *completely randomized experiment***, I, J individuals are randomly split into I treatment groups, with J individuals per group.
- But **heterogeneity** among individuals can inflate the random variation in the observed responses, making it harder to detect treatment effects.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Example

A study investigated the productivity of secretaries with different word processing programs. The study design called for giving an identical task to **nine** secretaries, allocated to **three** treatment groups.

Group 1 used a primarily **menu-driven** program. **Group 2** used a **command-driven** program and **Group 3** used a **mixture** of both approaches.

The time (in minutes) taken to complete the task was recorded.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

The secretaries had **different** levels of experience, typing speed, and computer skills.

If a ***completely randomized*** one-factor experiment was carried out, this **heterogeneity** would contribute to the **random variation** in completion times **within** each group.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Factor: Word Processing Program

Menu Driven	Command Driven	Mixture
13	14	11
10	12	8
8	9	7

Some of the observed variation within treatment groups is due to differences in experience levels.

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- In a **randomized block experiment**, the IJ individuals are first divided into J groups of I individuals per group, called **blocks**, that are **homogeneous** with respect to a so-called **blocking variable** that's believed to contribute to variation in observed responses.

Then, separately for each block, the I **individuals within the block** are **randomized** to the I **treatments**.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Example

For the secretary productivity study using a **randomized block experiment**, the **nine** secretaries are first split into **three blocks** (groups) of **three** secretaries each based on **experience level** (less than 1 year, 1 - 5 years, and more than 5 years).

Then, **within each block**, the **three** secretaries are **randomly assigned** to the three **word processing programs**.

The data are on the next slide.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Factor: Word Processing Program

		Menu Driven	Command Driven	Mixture
Blocks: Experience Level	< 1 Year	13	14	11
	1 - 5 Years	10	12	8
	> 5 Years	8	9	7

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Two-Factor ANOVA with $K = 1$ (Cont'd)

- In **randomized block experiments**:
 - The effects of the treatments are of major interest to the experimenter.
 - The effects of the blocking variable are generally not of interest.
- The analysis is carried out **exactly** as if the **blocking variable** was a **second factor** in the experiment.

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Example

For the study of secretary productivity using the *randomized blocks design*, the ANOVA table is below.

Source of Variation	df	Sum of Squares	Mean Square	f	P-value
Blocks (Experience)	2	32.89	16.444	59.2	0.00107
Treatments (Program)	2	13.56	6.778	24.4	0.00574
Error	4	1.11	0.278		
Total	8	47.56			

The **word processing program** has an **statistically significant effect** on the time to complete the task.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Notes

- A **randomized block experiment** explicitly **models** the **blocking variable** as a source of **deterministic** (non-random) variation in the data, thereby eliminating it as a contributor to **random variation**.

The goal is to **gain power** for detecting a **treatment effect**.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

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Example

In the secretary productivity study, if **blocking wasn't** used the data would be as shown below.

Factor: Word Processing Program

Menu Driven	Command Driven	Mixture
13	14	11
10	12	8
8	9	7

Some of the **random variation** within groups is due to differences in secretaries' experience levels.

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Two-Factor ANOVA with $K = 1$ (Cont'd)

Notes

The **one-factor ANOVA table** is below.

Source of Variation	df	Sum of Squares	Mean Square	f	P-value
Treatments (Program)	2	13.56	6.778	1.20	0.3650
Error	6	34.00	5.667		
Total	8	47.56			

The **SSE** here is much **larger** than when **blocking was** used.

(In fact, the **SSE here** is the the **SSE** for the **blocked model plus** the **SSA** for that model.)

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The **larger SSE** here leads to a **larger MSE**, **smaller F** value, and **non-significant** treatment (program) **effect**.

- **Comment:** Although blocking leads to a **smaller SSE**, it also leads to **fewer df** for **SSE** ($IJ - I - J + 1$ compared to $I(J - 1)$).

Thus blocking **can** lead to a **larger MSE** if the **reduction in SSE** is **small** relative to the decrease in **df**.

In this case, there's no advantage to blocking.

- **Comment:** A **matched pairs** study is a **randomized block experiment** in which there are **two treatment groups** and each **pair** is a **block**.

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